Validation of the positive and negative affect scale in the Chilean population and its application on migrant people

Héctor Arancibia Martini^a

^a Facultad de Humanidades y Educación, Universidad de Valparaíso, Valparaíso, Chile

*Corresponding author hector.arancibia@uv.cl

Citation Arancibia Martini H. Validation of the positive and negative affect scale in the Chilean population and its application in migrant people. *Medwave* 2019;19(1):e7579

Doi 10.5867/medwave.2019.01.7579

Submission date 20/12/2018 Acceptance date 28/1/2019 Publication date 22/2/2019

Origin not commissioned

Type of review reviewed by two external peer reviewers, double-blind

Key Words positive and negative affect scale, validation, migrant's well-being, hedonic well-being, feeling of happiness, life satisfaction

rison with migrants' people were evaluated.

Results

The reliability index was $\alpha = 0.92$ for the positive affect sub-scale and $\alpha = 0.75$ for the negative affect sub-scale. The exploratory factor analysis reproduced in the same way the factorial loadings proposed in the original instrument with saturations > 0.35 per item. Colombian and Venezuelan migrants show greater expression of positive affect than Chileans, while this difference is not observed when comparing Chileans with Peruvians.

Conclusion

The positive and negative affect scale has adequate psychometric properties to measure the expression of affect (hedonic well-being) in the Chilean population.

Abstract

Introduction

Although the positive and negative affect scale proposed by Mroczek and Kolarz has not been as widely used as other instruments, the adequate psychometric properties reported and the simplicity of the application, make it an adequate instrument to measure the hedonic component of subjective well-being. Thus, the scale validation in Chile will allow researchers to have a short and reliable instrument available in order to measure subjective well-being.

Objective

The initial objective of this study was to in Chile the "Positive and negative affect scale" proposed by Mroczek and Kolarz.

Methods

A non-probabilistic sample of 155 Chileans and a sample of 308 adult immigrants residing in Chile from Colombia, Peru and Venezuela were recruited. The sample of Chileans was used to validate the scale. The construct validity (Cronbach's α and exploratory factor analysis) and the concurrent validity through bivariate correlations with other measures were estimated. Finally, the differences in affect expression by Chileans in compa-



Main messages

- The validation in Chile of the scale of positive and negative affect proposed by Mroczek and Kolarz is presented.
- The validation was carried out employing statistical analysis (a de Cronbach, exploratory factorial analysis and bivariate correlations) in non-probability samples of Chilean and migrant populations. The differences in the expression of affection of the Chileans in comparison with that of the migrants were also valued.
- The generalization of the results is limited by the non-probability sample used in the study.
- The scale of positive and negative affection presents psychometric properties adequate to measure the expression of hedonic well-being in the Chilean population.

Introduction

The field of well-being could fall under the umbrella of two big traditions: one which focuses on human potential (eudaimonic well-being) and the second which centralizes on happiness (hedonic well-being)¹. The latter of these approaches is linked to subjective well-being and, according to evidence, is made up of three components: the expression of negative affect, the expression of positive affect and life satisfaction². Life satisfaction accounts for the distance perceived by people in relation to their aspirations, thus it is linked more to cognitive components. On the other hand, negative and positive expressions of affect are related to more immediate experiences, contingent or situational, that a person lives. The latter would not implicate instability in the expression of affect produced by the influence of emergent events given that, as Headey and Wearing³ suggest, after experiencing emergent events, people tend to adapt themselves to the situations, returning to the initial point of balance from before the event.

The balance of the expression of positive and negative affect is, from Bradburn's⁴ perspective, that which accounts for the degree of happiness perceived by people. Specifically, positive affect represents the dimension of positive emotionality which is expressed through joy, a good mood, from the sensation of plenty and tranquility in life. Conversely, the expression of negative affect represents emotional expression characterized by desperation, sadness, a sense of low social value and unease, among other aspects⁵.

As for the structure of the scale of positive and negative affect, although both expressions of affect (the negative and the positive) are related, they are represented by two different conceptual dimensions. It is for this reason that in the literature it is suggested that they should be measured separately⁶. Therefore, according to Bradburn's initial proposal, different studies have reported the distinction between both dimensions with moderate to high inverse correlations between them^{7,8,9,10}.

In relation to the explanatory variables of subjective well-being, these can be linked with genetic aspects, personality traits and social-cultural variables. Tellegen and collaborators¹¹ investigated the genetic component of happiness by following different pairs of twins who were temporarily separated, both monozygotic and dizygotic, estimating that 40% of the variability in positive affect and 55% in negative affect could be predicted by genetic variation. Regarding personality traits, those who tend to be extraverted present moderate to high correlations in the expression of positive affect¹².

The way in which people think about the world and their own circumstances also influences the expression of affect Those who have a greater disposition to recognize the positive aspects of life, that have realistic expectations regarding control of external situations, who are confident and cautiously optimistic; present greater expressions of positive affect¹³. The same occurs when these people develop goals and structure clear and feasible life projects. They tend to be happier, enthusiastic and with a greater level of vital energy.

Cultural context, also is described as one of the variables with more influence when explaining subjective well-being. In this sense, it is of special interest to differentiate individualist sociocultural contexts from collective ones. In the former, personal thoughts, options and feelings are more highly valued. This explains why self-esteem and the coherence of acts based on personal values correlate more strongly with the expression of positive affect in individualist societies than in collective ones. On the other hand, in collective societies, people are more willing to disregard themselves in favor of group interests and their expression of affect considers the link with others^{14,15}. In more individualist societies, people, when questioning their level of life satisfaction, generally make connections with their own feelings, unlike people in collective societies who, when faced with the same question, have in mind their family, friends and peers social evaluations and usually refer to social norms¹⁶.



In relation to the present study, the aim is to validate the Positive and Negative affect Scale proposed initially by Mroczek and Kolarz⁶ in the Chilean population in order apply then on resident migrant populations in Chile (Colombians, Peruvians and Venezuelans). The aim is to compare the results obtained by the sample of Chileans with the migrants to explore differences in the expression of positive and negative affect between the groups.

Ethical Aspects

This investigation was reviewed and approved by the Scientific Ethics Committee from the University of Valparaiso (Comité de Ética Científica de la Universidad de Valparaíso), which is accredited by the Ministry of Health (*Ministerio de Salud*) in Chile. The participants' information was incorporated anonymously to a database of the statistical program SPSS v.22.

Methodology

Participants and Procedure

The participants were given a survey after having read and signed the informed consent agreement, which guaranteed, among other aspects, the anonymity of the participants. The gathering of data was carried out by a team of professional survey-givers who applied the surveys either in public spaces or in the houses of the participants. They explained to the participants that the study was part of a research project which tried to account for different aspects of life between resident people in Chile.

The non-probabilistic sampling was gathered in two regions of Chile: The Valparaíso and the Metropolitan Region. For the purpose of scale-validation, 155 participants of Chilean nationality whose ages fluctuated between 18 and 78 (SD = 14.85) were conscripted. The sampling was designed by looking for a homogenous distribution in gender variables (n women = 83 [SD= 14.89]; n men = 72 [SD= 14.09]) and the level of education (n technical or less= 47.1%; n incomplete university education or superior= 52.9%). The validated instrument was subsequently applied to a sample of migrant participants of Venezuelan, Colombian and Peruvian nationalities residing in the same regions as the Chilean participants.



Table 1.	Socio-o	lemograp	hic variat	oles in	the	sample.
----------	---------	----------	------------	---------	-----	---------

	Chilear	ns	Colombians		Peruvian	S	Venezuelans	
Variable	ħ	%	п	%	п	%	n	%
Gender								
Female	83	53.5	60	55	58	47,9	70	53
Male	72	46.5	49	45	63	52.1	62	47
Age								
18-29	46	29.7	49	43.1	54	44.6	56	42.4
30-49	61	39.3	52	47.7	53	43.8	66	50
50-69	42	27.1	9	8.3	13	10.8	7	5.3
\geq 70	6	3.9	1	0.9	1	0.80	3	2.4
Education								
Secondary or less	39	25.2	40	36.7	54	44.6	33	25.1
Incomplete Technical	9	5.8	2	1.8	12	9.90	6	4.50
Complete Technical	25	16.1	27	24.8	29	24.0	34	25.8
Incomplete University	23	14.8	6	5.5	11	9.10	12	9.10
Complete University	59	38.1	22	20.2	15	12.4	63	47.7
Occupation								
Head of household	12	7.7	1	0.9	7	5.8	3	2.3
Student	15	9.7	3	2.8	2	1.7	0	0.0
Paid employment	122	78.7	103	94.5	111	91.7	126	95.5
Retired/pensioner	4	2.6	2	1.8	0	0	0	0.0
Unemployed	2	1.3	0	0	0	0	3	2.3
Marital/Civil status								
Single	75	48.4	56	51.4	66	54.5	72	54.5
Married/partnership	63	40.6	44	40.4	48	39.7	48	36.4
Widowed	4	2.6	1	0.9	1	0.8	1	0.8
Separated/divorced	13	8.4	7	6.4	6	5.0	8	6.1
Region of residency								
Metropolitan	78	50.3	53	48.6	71	58.7	46	34.8
Valparaiso	77	49.7	56	51.4	50	41.3	86	65.2
Religious practice								
None	71	45.8	18	16.5	19	15.7	24	18.2
A little	28	18.1	39	35.8	38	31.4	39	29.5
Somewhat	31	20	33	30.3	42	34.7	39	29.5
Considerable	18	11.6	15	13.8	16	13.2	20	15.2
Very considerable	5	3.2	2	1.8	4	3.3	7	5.3

Source: prepared by the author from the results of the study.

Instrument

The Positive and Negative Affect Scale. The validated Spanish version by Diaz, Blanco and Durán was used with one answer option added to avoid the preference of indecision. In this way the participants were asked to answer to a six-options Likert scale. These were: (1) never; (2) only once; (3) sometimes; (4) many times; (5) almost always; (6) always. Higher scores in each of the sub-scales indicated greater expression of that affect, so it was expected that the correlation between both scales would be medium to high with the negative coefficient. According to the results obtained by Mroczek and Kolarz the reliability coefficient α Cronbach was α =0.87 for the sub-scale of negative affect and α =.91 for the sub-scale of positive affect.

To evaluate the validity of the construct, the Cronbach's α from the full scale and of the two subscales was estimated with the Chilean sample. Subsequently the pertinence of the data was evaluated

to carry out an exploratory factor analysis (KMO and X^2 Bartlett). The eigenvalues of the matrix of correlations were estimated and the scree plot was presented in order to make the decision about the factors to extract.

After, the factor loadings were estimated for the respective factors. To evaluate the concurrent validity, bivariate correlations were done between the scores of the measurements. To achieve this, the scores in the scale of positive and negative affect were correlated with the measurement of the "Feeling of Happiness" proposed by Wold¹⁷ and with the general measurement of "Life Satisfaction" proposed by Keyes, Shmotkin & Ryff¹⁸. After conducting the estimations with the Chileans sample data, Student's T-test for independent samples were estimated to evaluate the existence of differences between the means of Chilean sample and that of the Peruvian, Colombian and Venezuelan migrants.

With the objective of predicting the scores in the dependent variables of positive affect and negative affect, a linear regression was estimated considering as independent variables gender, age, education and religious practice.

Results

According to Nunnally Berstein's criteria¹⁹ the full scale of positive and negative affect (12 items) presented a high reliability ($\alpha =$ 0.89) similar to the subscale of positive affect ($\alpha = 0.92$). On the other hand, the negative affect subscale presented an acceptable reliability rating ($\alpha = 0.75$). The correlation item-test for the entire scale (12 items) ranged from 0.256 to 0.775.

erain the itematis.removed and t

Table 2. Item-test Correlation Coefficients, Cronbach's

Itom	Item-test	Cronbach's α	Cronbach's	
Item	correlation	without the item	α	
1. So sad that I could never cheer up	0.605	0.886		
2. Nervous	0.539	0.889		
3. Uneasy	0.608	0.886		
4. Hopeless	0.542	0.889		
5. Everything is a huge effort	0.440	0.897		
6. Not valued	0.256	0.903		
7. Нарру	0.650	0.883		
8. In a good mood	0.736	0.879		
9. Very happy	0.700	0.880		
10. Calm and peaceful	0.713	0.880		
11. Satisfied	0.773	0.876		
12. Full of life	0.775	0.875		
Full scale positive and negative affect			0.89	
Negative affect subscale			0.75	
Positive affect subscale			0.92	

Note: The scores from items 1 to 6 of the negative affect subscale have been inverted.

Source: Prepared by the author from the results of the study.

Exploratory Factor Analysis

The analyses for the sample demonstrated adequate levels of adjustment for the factor analysis, where KMO = 0.894 and X^2 Bartlett (66; N=155) = 980.50; p< 0.001. The exploratory factor analysis was completed with the 12 items from the scale. The method of extraction used was the maximum likelihood estimation with direct Oblimin rotation. The results of the parallel analysis (Table 3) and the scree plot (Figure 1), suggest the extraction of two factors, compatible with the structure of the original scale.

Factor	Eigenvalue	Variance %
1	5.82*	48.54
2	1.49*	12.44
3	0.96	8.04
4	0.84	6.97
5	0.61	5.09
6	0.54	4.50
7	0.41	3.38
8	0.35	2.94
9	0.29	2.40
10	0.26	2.18
11	0.23	1.90
12	0.20	1.64

Tabla 3. Eigenvalues of the matrix of correlations between the 12 items.

*: Eigenvalues > 1.

Source: Prepared by the author from the results of the study.





On Table 4 are informed the items factor loadings of the scale. The results indicates that both factors grouped the corresponding items in a clear and differentiated way from one subscale to the other (high factorial saturations > 0.35), replicating the results in an identical way to the factorial structure proposed originally by Mroczek and Kolarz. The inter-factor correlation between the sub-scales was medium and negative (r = -0.528), which was expected given that one of them measures the expression of negative affect and the other of positive.

Source: Prepared by the author from the results of the study.

Table 4. Factor loadings from the items of the sub-scales and communalities

	Fa	ctor	h²	Sub-scale	
Item	Ι	II			
1. So sad that I could never cheer up		0.541	0.400		
2. Nervous		0.782	0.512		
3. Uneasy		0.820	0.567	N	
4. Hopeless		0.527	0.389	Negative affect	
5. Everything is a huge effort		0.462	0.327		
6. Not valued		0.351	0.215		
7. Нарру	0.831		0.659		
8. In a good mood	0.826		0.667		
9. Very happy	0.867	0.867 0.696 p		Desision offers	
10. Calm and peaceful	0.751		0.626	Positive affect	
11. Satisfied	0.768		0.675		
12. Full of life	0.843		0.693		
Inter-factor correlation	-0.528				

Source: Prepared by the author from the results of the study.

Concurrent validity

In table 5 the bivariate correlations are included among the full scale, the sub-scales, the measurement of "feeling of happiness" and the measurement of "life satisfaction". All of the correlations were high and statistically significant (p < 0.001). Given that the scores of the items on the full scale of positive and negative affect were inverted, a high and negative correlation was expected with the subscale of negative affect (r = -0.854; p < 0.001). According to what can be expected, the correlation between the full scale of positive affect and the measurement of "feeling of happiness" was

medium-high (r = 0.632; p < 0.001) similar to the measurement of "life satisfaction" (r = 0.558; p < 0.001).

The negative affect sub-scale correlated inversely with the measurement of "feeling of happiness" (r = -0.502; p < 0.001), which indicates that the people who reported greater expressions of negative affect also shared lower levels of subjective happiness. Likewise, those who reported greater levels of negative affect also presented lower levels of "life satisfaction" (r = -0.364; p < 0.001).

Table 5. Correlations and descriptive statistics.

	1	2	3	4	5
1. Full scale		-0.854	0.921	0.632	0.558
2. Negative affect subscale			-0.584	-0.502	-0.364
3. Positive affect subscale				0.608	0.592
4. Feeling of happiness					0.532
5. Life satisfaction					
Μ	54.87	12.48	25.35	2.97	7.69
DT	9.27	4.44	5.91	0.61	1.46

Note: All of the correlations were significant (p<0.001)

Source: Prepared by the author from the results of the study.

In executing a backwards method elimination with the sample of Chilean participants, where the score of the positive affect scale was included as a dependent variable and gender, age, education and religious practice as independent variables; the variables are ruled out (model 5 only includes the constant), which allows one to interpret that the independent variables incorporated in the model do not predict the score of the positive affect scale.

In the case of the negative affect scale (dependent variable), gender, age, education and religious practice were considered as independent variables. The final model that is adjusted to the data only includes the level of education variable (F (6.285.) = 1.9; p < 0.05). Therefore, one can conclude that this independent variable explains the changes in the scores of negative affect, ruling out the effect of the gender, age and religious practice. The typified regression equation that allows for the explanation and prediction of the negative affect scores is: z negative affect = -0.641^* (Zeducation). In having a negative value coefficient, it stands to reason that a greater level of studies indicates a lesser expression of negative affect.

Comparison of the mean between the samples

As a strategy to establish the scale's criterion validity, diverse mean contrasts were completed (Table 6). One of them was the contrast in the scores of the sample of Chileans and Colombians in the Negative Affect sub-scale. The difference was statistically significant, which allows for the affirmation that the Chilean sample express greater feelings of negative affect than the Colombians (M Chileans = 12.48 > M Colombians = 11.44; p < 0.05). This difference in the expression of negative affect is not observed when between Chileans and Venezuelans, nor between Chileans and Peruvians (p > 0.05).

The results of the mean contrast in the positive affect subscale accounts for significant differences in the expression of positive affect between Chileans and Colombians (*M* Chileans = 25.35 < M Colombians = 27.24; p < 0.05) and between Chileans and Venezuelans (*M* Chileans = 25.35 < M Venezuelans = 27.09; p < 0.05). The contrast of averages does not allow for the conclusion that there are significant differences between expression of positive affect by Chileans in regards to the expression of this affect by the Peruvians (*M* Chileans = 25.35 < M Peruvians = 25.87; p > 0.05).

Subscale	Group	Ν	М	SD	t (95% IC)	d.f.	d^{**}
	Chileans	151	12.48	4.44	-1.99*	254	0.22
	Colombians	105	11.44	4.72	(-2.080.02)		
Norseine effect	Chileans	151	12.48	4.44	0.030*	275	0.00
Negative affect	Venezuelans	126	12.50	4.72	(-1.07 - 1.11)		
	Chileans	151	12.48	4.44	-1.661	268	
	Peruvians	119	11.64	3.78	(1.85 - 1.60)		
	Chileans	153	25.35	5.91	2.79*	258	0.34
	Colombians	107	27.24	4.48	(0.56 - 3.22)		
De l'aliana affra at	Chileans	153	25.35	5.91	2.58*	275	0.30
Positive affect	Venezuelans	124	27.09	5.12	(0.41 - 3.06)		
	Chileans	153	25.35	5.91	0.783^{1}	272	
	Peruvians	121	25.87	4.96	(779 - 1.81)		

Table 6. Comparison of the scores obtained by subscale between the sample of Chileans with the Colombian, Venezuelan and Peruvian migrants.

 $^{1}p > 0.05$; $^{*}p < 0.05$. The effect size (Cohen's d^{**}) was calculated in the average comparisons considered statistically significant (p < 0.05).

Source: Prepared by the author from the results of the study.



Discusion

The purpose of this study was to validate in a chilean sample the Positive and Negative Affect Scale, originally proposed by Mroczek and Kolarz. Once the scale was validated, a comparison was made between the results obtained by the Chilean sample and those obtained by the Colombians, Venezuelans and Peruvians who reside in Chile.

The six items sub-scale (those of positive and negative affect) presented an excellent item-test correlation and adequate levels of reliability. The full scale and each of the sub-scales, showed excellent internal consistency, comparable with the original version. Through the exploratory factor analysis, the bifactor solution was the one that best fit the data. The distribution of the items, according to their factor loadings, was identical to the original structure. As one sub-scale measured the expression of positive affect and the other of negative, the correlation between both was moderated with a negative coefficient. These results also correspond with those obtained by other authors who have used the scale in different contexts^{17,19}. In turn, the correlations of medium and significant intensity (p < 0.001) between the sub-scale of positive affect and the measurements of "feeling of happiness" and "life satisfaction", like those obtained by the negative affect, but with a negative coefficient, show evidence of the concurrent validity of both sub-scales.

Comparing the scores between the samples, the Colombians obtained a greater level of positive affect expression and less of the negative when compared to the Chileans. These results are compatible with the results from the World Database on Happiness, where the Colombians are in first place among the 90 countries surveyed²⁰.

With these results it stands to reason that the level of happiness of the Colombians cannot be explained solely within their national context, but is a condition that is also maintained outside of it, despite the difficulties of a migratory experience, such as the one are currently carrying out in Chile.

Similarly, in Spain, an investigation with Colombian, Bolivian, Romanian, Sub-Saharan African and Moroccan migrants showed inferior levels of mental health than those of the Spanish, with the exception of the Colombians who obtained indicators similar to those of the native population²¹. In that regard, it is widely described as happiness, and thus the expression of it promotes positive indicators of mental and physical health, fosters positive social relationships, increases a person's vital energy and reduces stress and the rates of accidents and suicide among other aspects.

In general, the Colombian's high levels of happiness cannot be explained by objective conditions such as income level, country of residency, transportation, among others. Previous research have pointed out important variables to explain the happiness of Colombian people that are linked closely to everyday life, such as family, their affectionate relationships, work environment and some individual conditions such as self-efficacy²³. The data also reported that Venezuelans show a greater expression of positive affect in comparison with Chileans. This result could be explained by the aforementioned reasons with regards to the Colombian sample; one could conceive that, due to the close geographical proximity of people with these two nationalities, there is a common foundation in which the appreciation of personal relationships as well self-efficacy are shared.

One area where there were not significant differences, neither in the expression of positive affect nor in the negative, was between the sample of Chileans and Peruvians. Those results support the low impact of the "country of residence" variable to account for people's subjective well-being. In the Chilean's case, the lower levels of expression of positive affect and the higher expression of the negative could be explained by the individualistic society that they are a part of, characterized by high levels of competitiveness, significant stress levels and a reduced social support network. In this sense, a reasonable explanation of the results between Chileans and Peruvians may be due to an assimilatory time effect by the Peruvian migrants, given that the Peruvian migration has a long history in Chile unlike that of Venezuelans and Colombians.

This investigation has some limitations. The sample is not a probability sample so caution should be taken when generalizing the results. In order to have a greater understanding of the differences in scores between the samples in the expressions of affect subscales, other variables should be incorporated into the analysis to enrich the interpretation of the results. In this sense, it would be relevant to consider the impact of the effects of moderating and the mediating variables in the explanation of the expression of positive and negative affect.

Conclusion

The positive and negative affect scale present psychometric properties that are adequate for measuring the expression of affect (hedonic well-being) in the Chilean population.

Colombian and Venezuelan migrants show greater expression of positive affect than the Chileans, whereas this difference is not observed when comparing Chileans with Peruvians.

In future research, measuring the life satisfaction variable together with the expression of positive and negative affect is suggested in order to verify the factorial structure of subjective well-being in the Chilean population.

Finally, it is important to mention possible new lines of investigation that arise from this investigation, such as the execution of a validation exercise with a confirmatory analysis in a representative sample or the exploration for proof of validity and reliability in the migrant population.

Notes

Declaration of Funding

This study was financed by the *Comisión Nacional de Ciencia y Tecnología (CONICYT)*, through the Postdoctoral *FONDECYT* N° 3170187.

Declaration of ethical aspects and conflicts of interest

The author declares that this investigation was reviewed and approved by the *Comité Ético Científico* from the *Universidad de Valparaíso*, following the *Acta de Aprobación Bioética* CEC142-17 on the date of April 11, 2017.

Declaration of conflict of interests

The author declares that there are no potential conflicts of interest in regards to the investigation, authorship, and/or with the publication of this article.

References

- Ryan RM, Deci EL. On happiness and human potentials: a review of research on hedonic and eudaimonic well-being. Annu Rev Psychol. 2001;52:141-66. | CrossRef | PubMed |
- Diener E, Lucas RE, Scollon CN. Beyond the hedonic treadmill: revising the adaptation theory of well-being. Am Psychol. 2006 May-Jun;61(4):305-14. | PubMed |
- Headey B, Wearing. A. Understanding happiness: A theory of subjective well-being. Melbourne: Longman Cheshire; 1992. [on line]
 Link |
- 4. Bradburn, NM. (). The structure of psychological well-being. Oxford, England: Aldine; 1969. | Link |
- Clark LA, Watson D, Mineka S. Temperament, personality, and the mood and anxiety disorders. J Abnorm Psychol. 1994 Feb;103(1):103-16. | PubMed |
- Mroczek DK, Kolarz CM. The effect of age on positive and negative affect: a developmental perspective on happiness. J Pers Soc Psychol. 1998 Nov;75(5):1333-49. | PubMed |
- Cacioppo JT, Gardner WL, Berntson GG. The affect system has parallel and integrative processing components: Form follows function. Journal of personality and Social Psychology, 1999;76(5):839-855. | CrossRef |
- 8. Diener E, Emmons RA. The independence of positive and negative affect. J Pers Soc Psychol. 1984 Nov;47(5):1105-17. | PubMed |
- Diener E, Smith H, Fujita F. The personality structure of affect. Journal of Personality and Social Psychology. 1995;(69):130-141. | Link |

- Russel JA. A circumflex model of affect. Journal of Personality Social Psychology. 1980;(39):1161-1178.
- Tellegen A, Lykken DT, Bouchard TJ Jr, Wilcox KJ, Segal NL, Rich S. Personality similarity in twins reared apart and together. J Pers Soc Psychol. 1988 Jun;54(6):1031-9. | PubMed |
- 12. Lucas RE, Fujita F. Factors influencing the relation between extraversion and pleasant affect. J Pers Soc Psychol. 2000 Dec;79(6):1039-56. | PubMed |
- Diener E, Lucas R, Oishi S. Subjective Well-being. The Science of Happiness and Life Satisfaction. En: Handbook of positive psychology. England: Oxford University Press;2002:63-73.
- Diener E. Subjective well-being. The science of happiness and a proposal for a national index. Am Psychol. 2000 Jan;55(1):34-43. | CrossRef | PubMed |
- Suh E, Diener E, Oishi S, Triandis HC. The shifting basis of life satisfaction judgments across cultures: Emotions versus norms. Journal of Personality and Social Psychology. 1998;74(2):482-493. | CrossRef |
- Díaz D, Blanco A, Durán MM. La estructura del bienestar: el encuentro empírico de tres tradiciones. Revista de Psicología Social. 2011;26(3):357-372. | CrossRef |
- 17. Wold B. HBSC: A WHO Cross-National Survey. Resource Package of Questions 1993-1994. Norway: University of Bergen; 1995.
- Keyes CL, Shmotkin D, Ryff CD. Optimizing well-being: the empirical encounter of two traditions. J Pers Soc Psychol. 2002 Jun;82(6):1007-22. | PubMed |
- 19. Nunnally J, Bernstein I. Psychometry theory (3 ed.). New York: McGraw-Hill; 1994. | Link |
- 20. Veenhoven R. Average happiness in 90 nations 1990-2000: How much people enjoy their life-as-a-whole on scale 0 to 10 (No. WDoH RankReport2004/1b). 2004
- 21. Sevillano V, Basabe N, Bobowik M, Aierdi X. Health-related quality of life, ethnicity and perceived discrimination among immigrants and natives in Spain. Ethn Health. 2014;19(2):178-97. | CrossRef | PubMed |
- 22. Lyubomirsky S, King L, Diener E. The benefits of frequent positive affect: does happiness lead to success? Psychol Bull. 2005 Nov;131(6):803-55. | PubMed |
- 23. Gómez V, De Posada CV, Barrera F, Cruz JE. Factores predictores de bienestar subjetivo en una muestra colombiana. Revista latinoamericana de Psicología. 2007;39(2):311-325. | Link |

Correspondence to Serrano 546 Valparaíso Región de Valparaíso Chile



Esta obra de Medwave está bajo una licencia Creative Commons Atribución-No Comercial 3.0 Unported. Esta licencia permite el uso, distribución y reproducción del artículo en cualquier medio, siempre y cuando se otorgue el crédito correspondiente al autor del artículo y al medio en que se publica, en este caso, Mewave.